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**Local economic effects of connecting to China's high-speed rail network:  
Evidence from spatial econometric models**

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## **Abstract**

China's high-speed rail (HSR) has quickly expanded to over 40,000 km of lines operating and another 10,000 km under construction. This is over 10-times longer than the networks in long-established HSR countries like France, Germany or Japan. While fewer than 100 county-level units had stations on the HSR network in the first years of operation, the eight years from 2012-19 saw almost 400 more county-level units connect to the HSR network. Effects on local economic activity from this substantial increase in connections to the HSR network remain contested. Some prior studies find either insignificant effects on local economic growth or even negative effects in peripheral regions. In light of this debate, we use spatial econometric models for a panel for almost 2500 county-level units to study effects of connecting to the HSR network. We especially concentrate on the 2012-19 period that has high quality night-time lights data to provide an alternative to GDP as an indicator of growth in local economic activity. Our spatial econometric models allow for spatial lags of the outcomes, of the covariates, and of the errors. We also address potential endogeneity of the HSR networks and connections, using an instrumental variables strategy. Across a range of specifications, we generally find that growth in local economic activity is lower following connection to the HSR network, with this effect especially apparent when using high quality night-time lights data for the 2012-19 period. Hence, expansion of the HSR network may not boost China's economic growth.

### **JEL Codes**

R12

### **Keywords**

High-speed rail  
infrastructure  
luminosity  
spatial spillovers  
China

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## I. Introduction

Transportation infrastructure is generally regarded as a key driver of regional economic growth and development (Ma, Chen, and Yang 2020; Egger, Loumeau, and Loumeau 2023). High-speed rail (HSR) is a significant transport innovation (Wang, Jiang, and Miao 2023; Li, Yu, and Ma 2024), with advantages of speed, efficiency, energy saving, safety, punctuality, and stability (Li et al. 2021; Wang et al. 2023; Liu, Diao, and Lu 2024). Since the first HSR route opened in Japan in 1964, more than 20 countries have commenced HSR operations (Wang et al. 2024). Although China was a latecomer to HSR construction, it has developed extremely fast (Mei, and Zhang, 2021). Starting with the opening of the Beijing-Tianjin intercity HSR in 2008, China developed the world's longest HSR network with the highest transportation density (Banerjee et al. 2020; Wang, Jiang, and Miao 2023; Tang et al. 2024). According to the State Council Information Office, China's HSR network exceeded 43,700 km by the end of November 2023, which ranks first in the world (Ren et al. 2024). In addition, China's HSR network is still expanding. The China National Railway Group Limited stated that China's railway network will eventually span approximately 200,000 kilometers, including about 70,000 kilometers of HSR, by 2035 (Li and Li, 2024).

A growing literature studies local economic development impacts of HSR networks. For example, Ahlfeldt and Feddersen (2018) find that HSR stimulated the economic vitality of cities in Germany. Carbo et al. (2019) studied economic effects of HSR in Spain from 1995 to 2014 using difference-in-differences (DID) and found positive effects on urban economies. Based on a Computable General Equilibrium (CGE) model, Kim and Yi (2019) found that HSR benefitted Gross Regional Product (GRP) growth in South Korea. Likewise, positive effects of HSR on economic activity are found in studies from Italy (Cascetta et al., 2020), France (Facchinetti-Mannone, 2019) and Japan (Miwa, Bhatt, and Kato, 2022).

For China, scholars have used various methods to study HSR impacts in different regions. The empirical evidence on impacts tends to be mixed. Most studies show significant positive economic effects on cities along HSR routes. For example, using DID, Diao (2018) and Wang et al. (2022) found HSR had a positive role in promoting provincial economies by increasing GDP and accessibility. With a CGE model, Yang et al. (2023) found that HSR could stimulate GDP growth and reduce regional disparities in eastern China. Relatedly, Jin et al. (2024) focus on accessibility and connectivity features of HSR, and the role of the HSR network in alleviating urban-rural income disparity. Li and Li (2024) find that the opening of HSR lines positively stimulates economic growth, primarily through market size.

However, some researchers have found that HSR may have negative economic effects. For example, Qin (2017) found that GDP of counties along China's HSR routes are reduced by 3%-5% on average, with this decline operating through decreased fixed asset investments. Using a DID model, Gao et al. (2020) estimate a long-lasting negative effect of HSR on GDP per capita within the Yangtze River Delta region, of approximately 10%, mainly due to the population reallocation. Liu et al. (2022) found that large scale inter-regional HSR reduced labor productivity by almost 13% and the productivity gap between central and peripheral cities in the network widened after connecting to the HSR network.

Despite this growing body of empirical research, a key aspect of the potential impact on local economic activity from connecting to the HSR network remains largely unexamined. Most previous studies only explore the direct impact of HSR opening on regional economic development. This may be insufficient given that transport infrastructure, and especially HSR, shortens travel time and improves inter-city accessibility, and thus contributes to the flow of capital, ideas, population, and other factors between cities (Shi, and Wang, 2024). Consequently, the effects induced by transport infrastructure are likely to extend to the neighbouring areas. These spillovers are often ignored in applied studies, which use impact evaluation frameworks that treat the spatial units as independent of their surrounding areas. Hence, the results from prior studies may not represent the complete impacts if these spatial spillover effects are not considered.

To overcome this gap in the literature, the current study uses a comprehensive set of spatial econometric models to examine the impacts that connecting to the HSR network has on local economic activity. We use a decade-long panel of almost 2500 county-level units in China to examine HSR impacts on GDP and on luminosity as indicators of local economic activity. The research framework not only considers the spillover effects of connecting to the HSR network but also allows for the endogenous placement of HSR routes. In comparison to much of the existing research, our focus is on economic activity at a more local level, but with our sample covering almost all of China. Second, previous studies often use the DID model to test the impacts of HSR operation but do not consider spatial factors such as spillovers. In contrast, spatial spillover effects of HSR can be captured with spatial econometric models. In our framework, we start with a very general spatial autoregressive model with spatial autoregressive errors (SARAR) that encompasses popular models like the spatial Durbin model, the spatial lag model, and the spatial error model. Lastly, by using an instrumental variables (IV) strategy, we can assess whether the endogenous placement issue interferes with

estimation of the impact of HSR on the regional economy. Our analysis not only advances the understanding of how HSR affects economic growth at the county level in China but also provides insight into the planning of future HSR networks.

The subsequent parts of the paper are structured as follows. Section 2 reviews the literature related to HSR and discusses the main contribution of the study. Section 3 describes the data used in this study and the econometric modelling framework, including the spatial econometric models, and the instrumental variables strategy used to allow for the endogenous placement of HSR. Section 4 presents the empirical results, using both county-level GDP and night-time lights as indicators of economic activity. Finally, Section 5 concludes the study, and discusses the limitations.

## **II. Literature review**

In this section, we review literature on the impact of high-speed rail (HSR) on local economic development in China. We focus especially on the empirical methods and the indicators of local economic activity. The reviewed studies cover a wide range of topical issues, including whether to take into account the spatial econometric models and the HSR endogenous placement issues.

Rail transport is a crucial mode of inter-region transportation in China, a large developing country with a widespread geographical distribution of natural resources and population (Ke et al. 2017). HSR has become the fastest-growing transportation mode in China due to the advantages of efficiency, punctuality, and stability (Xu, and Huang, 2019; Tian et al., 2021). As a result of the rapid development, research on the impacts of HSR networks has become a hot topic in China (Luan, Guo, and Liang, 2024). However, the impact of HSR on the regional economic development of peripheral cities along its line is still under debate. Researchers use a variety of spatial scales for these studies, and different conclusions concerning the promotion, inhibition, or insignificance of HSR on local economic growth have been reported.

Studies of the impact of HSR on regional economic growth (both within and beyond China) are mostly based on methods such as the synthetic control method (Albalade, Campos, and Jiménez, 2023), computable general equilibrium (CGE) model (Kim and Yi, 2019), and difference-in-differences (DID) model (Chen, Lv, and Zhang, 2024). For example, Yu et al. (2021) applied the synthetic control method and found that HSR raises a county's total GDP.

Using a dynamic CGE model, Chen (2019) found that HSR infrastructure development in China has generated a positive regional economic impact. Compared to the above methods, the DID model has been used more broadly to investigate HSR's impacts on socioeconomic conditions. For example, Chen et al. (2023), and Liang et al. (2020) used county-level panel data to study the economic impacts of HSR, which showed that the opening of HSR led to economic growth along the route. Wang et al. (2023) used GDP data to investigate the heterogeneous economic effects of HSR on prefecture-level cities. In addition to GDP, nighttime lights data are recognized as an effective proxy index related to human activities and urban economic development. Guo et al. (2023) adopted nighttime light data as a proxy for regional economic activity and found that HSR had a positive impact.

Considering that transport infrastructure in a region not only affects local economic growth but also economic activities in adjacent areas, some studies used spatial econometric models to examine the aggregate growth effect and potential spatial spillover effect of HSR. Liu, Tang, and Wang (2024) showed that the construction of HSR led to a notable increase in income through a positive spatial spillover effect. Based on the Spatial Durbin model (SDM), Sun et al. (2023) and Huang (2021) analyzed the direct and indirect effects of HSR on economic development. However, not all uses of these models find positive economic spillovers for regional economic development. Jin et al. (2020) found that the spillover effect is insignificant based on the SDM model.

The selection of high-speed rail stations is usually considered to be non-random and biased towards cities with better economic development. Given these endogeneity issues, many studies used an Instrumental variables (IV) strategy. To construct an IV for actual HSR connections, previous studies either referred to historical information or used the straight-line strategy. Historical transport routes were mainly located based on the original local geological conditions due to the then-technical restrictions, and the original geological conditions were highly related to the construction of transportation infrastructure, such as for post-cities. Hence, post-cities from the Ming Dynasty (1368—1644) were chosen as the instrumental variables for transportation infrastructure in the regression models of Guan, Chen, and Li (2023), and Wang, Jiang, and Miao (2023). Similar practices are followed by Chen, Cheng, and Zhang (2023), and Guo et al. (2021), who used China's historical transportation network in 1961 as the IV for the current HSR connections.

In contrast to the historical instruments, the straight-line strategy is based on connecting large central cities, using the shortest distance (which should also result in fastest journeys

notwithstanding potentially higher construction costs). The instrumental variable obtained by the minimum spanning tree should be highly relevant to the actual HSR line and thus is used to construct IVs for the actual transport connection. For example, Qiu, Liu, and Liao (2023), and An et al. (2022) constructed an IV by drawing straight lines among major cities, to predict the location of the current HSR connection. Shen, Li, and Li (2023), and Chong, Chen, and Qin (2019) used both the spatial econometric model and an IV strategy to explore the economic impacts of HSR, finding that the impacts were not significant.

To help summarize this large and growing literature on the impacts of HSR on regional economic activity in China, 30 recent studies are described classifying them along two dimensions: whether they allow for or ignore spatial spillovers and whether they treat the placement of the HSR networks as endogenous or exogenous (Table 1). Despite broad discussions about the effects of HSR connection on regional economic development, it can be seen that the spillover effects of HSR have received inadequate attention, with far more studies in Table 1 ignoring spillovers ( $n=23$ ) than allowing for spillovers ( $n=7$ ). The combination of allowing for spillovers and allowing for endogenous placement is very rare ( $n=2$ ). Our paper adds to this small literature because we consider both of these issues in our econometric models.

In summary, existing studies on the economic impact of high-speed rail have not yet reached consistent conclusions, perhaps due to the differences in their spatial scope and research methods. In terms of spatial scale, there is little systematic research focused on the county-level economic impacts. Most of the previous HSR impact studies are based on panel data from provinces or prefectural cities, while panel data from county-level areas are less commonly used. This paper aims to address these gaps.

Table.1 Selected studies analyzing HSR impacts

Authors	Year	Time period	Spatial units	Spatial approach	IV approaches	Indicator(s)	Result	Objective
<i>Panel A: Studies that assume exogenous placement and no spatial spillovers</i>								
Chen et al.	2023	2008-2017	328 counties	/	/	GDP	+ve	To assess impacts of HSR opening on the inter-county economic gap
Wang, Yu, and Zhang	2023	1999-2013	333 prefectures	/	/	GDP	+ve	To find the heterogeneous economic effects of distance to HSR
Wang et al.	2023	2007-2018	62 prefecture-level cities	/	/	GDP	+ve	To estimate direct impacts of opening HSR on regional economic development
Yu et al.	2021	2001-2017	54 counties in Hubei Province	/	/	GDP	+ve	To analyse impacts of HSR on economic growth
Guo et al.	2020	2002-2018	24 cities	/	/	Luminosity	+ve	To investigated the impact of the HSR on urban economic development using NTL
Li, Wu, and Zhao	2020	2001-2017	40 prefecture-level cities	/	/	GDP	+ve	To investigate the economic effect of HSR
Liang et al.	2020	2012-2017	632 counties	/	/	GDP	+ve	To explore impacts of HSR on economic growth along the route
Chen	2019	2002-2013		/	/	GDP	+ve	To access the regional economic impacts of HSR
<i>Panel B: Studies that assume exogenous placement but allow for potential spatial spillovers</i>								
Liu, Tang, and Wang	2024	2000-2018	286 prefecture-level cities	SDM	/	Income	+ve	To investigate the impact of HSR on urban residents' income
Sun et al.	2023	2003-2018	285 prefecture-level cities	SDM	/	GDP	+ve	To examine joint impacts of HSR on the urban economy
Huang	2021	2008-2018	281 cities	SDM	/	GDP	+ve	To discuss the spatial-temporal heterogeneity of the relationship between HSR and the urban economy

Yu	2021	2005-2013	47 prefecture-level cities	SDID	/	GDP	+ve	To investigate the influence of Beijing-Shanghai HSR on regional economy
Jin et al.	2020	2002-2016	285 cities	SDM	/	GDP	+ve	To examine the role of HSR in influencing economic growth and disparity

*Panel C: Studies that allow for endogenous placement but do not allow for potential spatial spillovers*

Chen, Cheng, and Zhang	2023	2003-2014	351 cities	/	1961 train station	GDP	+ve	To study the effects of HSR construction on firms' cross-regional development
Guan et al.	2023	2011-2019	41 cities in YRD region	/	Ming Dynasty post station	GDP	+ve	To test the impact of HSR on high-quality economic development
Qiu, Liu, and Liao	2023	2003-2019	284 prefecture-level cities	/	Straight lines	GDP	+ve	To investigate the impact of railway express on urban carbon emissions reduction
Wang, Jiang, and Miao	2023	2003-2019	277 prefecture-level cities	/	Courier roads in the Ming Dynasty	GDP	+ve	To examine HSR's impact on urban resilience
Wang, Zhou, and Guo	2023	2010-2018	34 prefecture level cities	/	Railway stations in 1993	Urban tourism incomes	+ve	To explore the spatial impact of HSR on tourism economy
Yang and Ma	2023	2001-2020	285 cities	/	Railway stations in 1912s	GDP	+ve	To study the impact of the border effect on HSR links
An et al.	2022	2005-2018	40 prefecture cities in YRD region	/	Straight lines	GDP	+ve	To investigate the effects of HSR on local economic development
Chen and Wang	2022	2003-2019	291 prefecture-level cities	/	Straight lines	GDP	+ve	To study the impact of HSR on the consumer service industry
Liu et al.	2022	2001-2015	236 prefecture-level cities	/	Straight lines	GDP	-ve	To explore effects of HSR on labor productivity
Wang, Wu, and Liu	2022	2001-2016	219 prefecture-level cities	/	Straight lines	GDP	+ve	To explore effects of HSR connection on urban innovation

Dong et al.	2021	2000-2020	180 towns	/	Straight lines	GDP	+ve	To explores the determinants of HSR new towns' economic growth
Guo et al.	2021	2004-2010	20 cities	/	1961 China railway map	GDP	+ve	To investigate the impact of HSR on regional water environment
Kuang, Liu, and Zhu	2021	2004-2017	141 cities	/	1962 China railway map	GDP	+ve	To study the relationship between HSR connection and firm performance
Gao et al.	2020	2006-2015	211 county-level units in YRD region	/	Straight lines	GDP	-ve	To investigate effects of connecting to an HSR line on the local economy
Yao et al.	2019	2010-2014	285 cities	/	China's historical railway network	GDP	+ve	To study the HSR impacts on economic growth

*Panel D: Studies that allow for endogenous placement and also allow for potential spatial spillovers*

Shen, Li, and Li	2023	2011-2020	229 cities	GS2SLS	The post cities in the Ming Dynasty	GDP	ns	To study the influence of transportation infrastructure on digital economy
Chong, Chen and Qin	2019	2008-2015	268 cities	SAR, SEM, SDM	Historical routes	GDP	+ve	To evaluate the economic benefits of HSR network from the perspective of connectivity improvement

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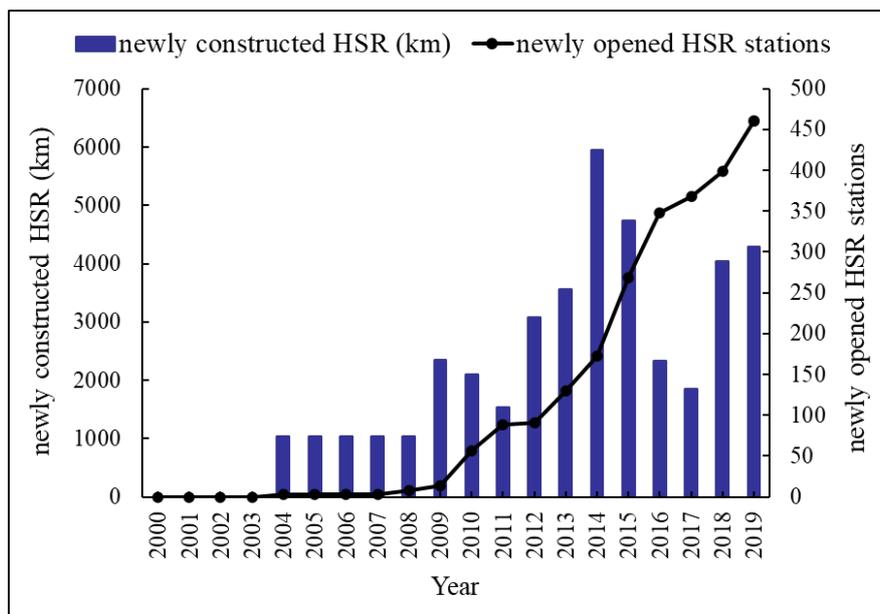
*Notes:* Abbreviations for spatial approach: GS2SLS (generalized spatial two stage least squares), SAR (spatial autoregressive model), SDM (spatial Durbin model), SEM (spatial error model), SDID (spatial difference-in-differences); +ve refers to positive effects, -ve refers to negative effects, ns refers to insignificant effects.

### III. Data and Methods

#### *Data Sources*

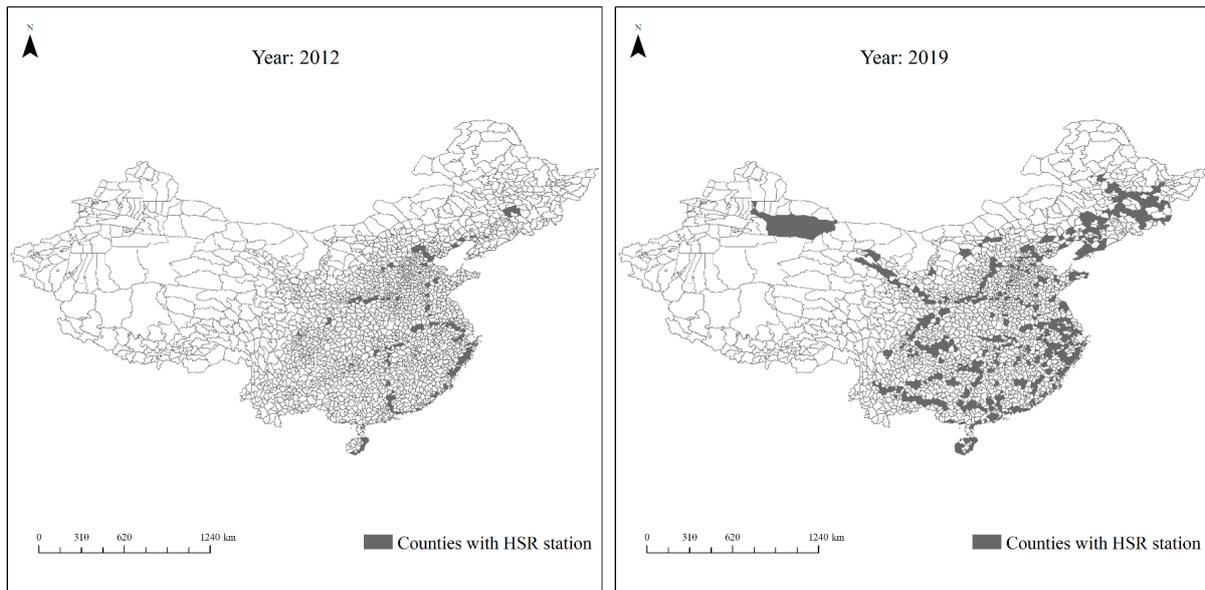
High-speed rail refers to high-speed trains with a design speed of over 250 km/h or passenger dedicated intercity lines with an average speed of 200 km/h or more. The HSR data were manually collected and include information such as the opening dates of each of China's HSR corridors in operation by the end of 2019, and the new length for each year. The sources for high-speed rail data include the National Railway Corporation ([www.china-railway.com.cn](http://www.china-railway.com.cn)) and the 12306 China Railway (<http://www.12306.cn>). For stations opening later than July 1 in a year, the opening date is set as the next year (the dummy variable  $HSR_{open} = 1$ ). The development of HSR in China began with the opening of the first HSR line at the end of 2008. Since then, the HSR network has experienced rapid expansion, particularly starting from 2012 (as shown in Fig. 1). Thus we adopt the research period from 2012 to 2019 (which also coincides with the availability of high quality luminosity data). During the research period, a total of 460 counties (out of 2342 counties in all of China) had HSR stations opened. The spatial distribution of HSR stations in China in 2012 and in 2019 is shown in Fig. 2. The expansion westwards, and to the northeast, and the increased density in central and southeast China is evident.

**Figure 1.** The growth of China's high-speed rail (HSR) network



Note: The data are sourced from the National Railway Administration of the people of China (<http://www.nra.gov.cn>) and National Bureau of Statistics ([www.stats.gov.cn](http://www.stats.gov.cn))

**Figure 2.** HSR stations in China as of 2012, and 2019



Note: The darkest areas, which take the value of 1 in the econometric models, are the counties and districts connected by HSR lines; the areas taking the value of 0 are counties and districts not connected by HSR lines.

To empirically analyze impacts we use two measures of local economic activity—GDP and luminosity. The GDP data were obtained from three main products of the National Bureau of Statistics (NBS): (i) annual editions of the China Statistical Yearbook (county-level) (in Chinese it is *Zhongguo Xianyu Tongji Nianjian (Xianshi Juan)*), (ii) annual editions of the China City Statistical Yearbook (known as *Zhongguo Chengshi Tongji Nianjian*), and (iii) annual editions of the Statistical Yearbook for each city or province (for example, the Beijing Statistical Yearbook) (NBS, various dates). Each edition reports on GDP the previous year, so we use the 2013 to 2020 editions to obtain annual GDP data from 2012 to 2019. Overall, we have a balanced panel of annual GDP data for each of  $n = 2342$  units at the 3rd level of the sub-national administrative hierarchy, where these units maintain a consistent spatial definition from 2012 to 2019.

The luminosity data we select to represent the level of regional economic growth are the annual composites from Defense Meteorological Satellite Program (DMSP) satellites and from the Visible Infrared Imaging Radiometer Suite (VIIRS) from the Suomi-NPP satellites. The DMSP annual composites from satellites F15 and F18 collectively cover each year from 2012 to 2019. The stable lights product provides 6-bit digital numbers (DN) from 0–63 for each 30 arc-second output pixel (approximate  $1.04 \text{ km} \times 1.04 \text{ km}$  at China’s latitude). Ephemeral lights, such as from fires and flaring, are removed and processing excludes (at pixel level) images for nights affected by clouds, moonlight, sunlight, and other glare. The VIIRS data we

used is the version 2.1 VIIRS Day/Night Band nighttime lights (VNL), which are available beginning in 2012. The VNL is produced from monthly cloud-free radiance averages, with initial filtering to remove extraneous features such as fires and aurora before the resulting rough annual composites are subjected to outlier removal procedures (Elvidge et al., 2021). The data are in units of nano Watts per square centimeter per steradian (nW/cm<sup>2</sup>/sr) reported on a 15 arc-second output grid (approximate 0.52 km × 0.52 km). We mainly use the average masked data product with background noise, biomass burning, and aurora zeroed out, which had the highest predictive power for GDP amongst all of the VNL data products in a prior study at the same county-level resolution that the current study uses (Zhang, and Gibson, 2022). Our constructed panel dataset contains annual observations for 2342 county-level units (which includes districts, county-level cities and counties) in China from 2012 to 2019. Table 2 shows the descriptive statistics.

Table 2. Descriptive statistics of variables

Variable	Sample size	Mean	Standard deviation	Minimum	Maximum
HSRopen	18736	0.119	0.324	0	1
IGDP	18736	4.883	1.284	0.262	10.508
IDMSP	18736	8.400	1.353	0	12.829
IVNL	18736	8.017	1.329	1.816	13.214

Notes: Statistics are for n=2342 spatial units, annually observed in a balanced panel from 2012 to 2019.

### *Estimation Framework*

Spatial econometric models let us examine the nature of possible spillovers, and are used for this purpose in many contexts (LeSage and Pace, 2009). A key aspect of these models aiding the study of spillovers is that possible interactions between spatial units are summarized with a  $N \times N$  spatial weights matrix,  $W$ . In this study we use a row normalized contiguity weights matrix that has values of one for neighbours and zero otherwise, with a diagonal of zeros because a spatial unit cannot neighbour itself. At the level of spatial disaggregation that we use, the average spatial unit in China has six neighbours.

In what follows, the proxy for economic output in spatial area  $i$  in year  $t$  is denoted as  $O_{it}$ , where the two proxy variables we use are log GDP in our main specification, and the log of the sum of night-time lights in our sensitivity analysis. The indicator for whether a spatial

unit has a HSR station is  $D_{it}$  the  $\mu_i$  are time-invariant fixed effects for each spatial unit, the  $\vartheta_t$  are year fixed effects, and  $e_{it}$  is a random error. By using the spatial weights matrix we can allow for spatial lags, which are averages of these variables over the neighbouring units.

Our starting point is a very general model, which is a spatial autoregressive model with spatial autoregressive errors (SARAR). This model allows for spatially lagged dependent variables, spatially lagged independent variables and spatially lagged errors:

$$O_{it} = \lambda W O_{it} + \beta_1 D_{it} + \beta_2 W D_{it} + \mu_i + \vartheta_t + \rho W u_{it} + e_{it} \quad (1)$$

The SARAR model allows for changes in an outcome variable in a given area to have effects on contemporaneous outcomes in other areas (via the autoregressive spatial lag of the dependent variable, if  $\lambda \neq 0$ ). It also allows changes in independent variables (such as getting connected to a HSR network) to affect not only own-area outcomes but also outcomes in neighbouring areas (if  $\beta_2 \neq 0$ ). The  $\rho W u_{it}$  term allows for spatial autocorrelation, where errors for a given area correlate ( $\rho$ ) with a weighted average of the errors from surrounding areas. Equation (1) nests a spatial Durbin model if  $\rho = 0$ , a spatial auto-regressive model (*aka* spatial lag model) where only the dependent variable is spatially lagged if  $\beta_2 = \rho = 0$ , a spatial error model where only the errors are spatially lagged (if  $\lambda = \beta_2 = 0$ ), and the most restrictive of all, which is an aspatial model with no spatial lags (if  $\lambda = \beta_2 = \rho = 0$ ). The aspatial model has been the approach underpinning many previous studies of HSR impacts on the regional economy (as shown in Table 1). The encompassing nature of equation (1) allows for a general-to-specific model selection strategy which appears to be more robust than the reverse simple-to-general selection strategy, especially if there are any anomalies in the Data Generating Process (Mur and Angulo, 2009).

An important feature of spatial econometric models is that lags of either the outcome variable or of independent variables (but not of errors) mean that total effects of changes in an independent variable—e.g. whether a county gets connected to a HSR network—may be quite different to what the regression coefficient on the dummy variable for being connected shows. Thus, while  $\hat{\beta}_1$  is the object of interest in the typical model without spatial lags, in the spatial models when either the spatial lags of outcomes or the spatial lags of independent variables are non-zero then  $\hat{\beta}_1$  does not capture the total effect of a change in the administrative status of a county. A useful decomposition of the more complex spatial relationships that occur relies on rewriting equation (1) in matrix notation (for simplicity, subscripts are dropped and fixed effects and error terms combined in  $v$  because the errors do not affect this decomposition) as:

$$O = (I - \lambda W)^{-1}(D\beta_1 + WD\beta_2) + (I - \lambda W)^{-1}v \quad (2)$$

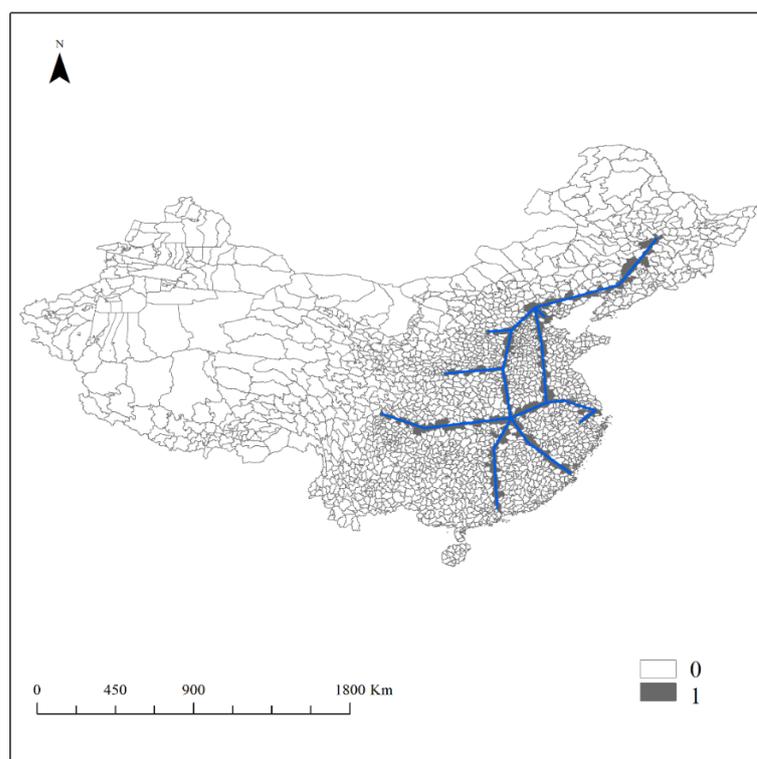
Following Elhorst (2012), the  $N \times N$  matrix of partial derivatives can be written (noting that diagonal elements of  $W$  are zero) as:

$$\frac{\partial O}{\partial D_k} = (I - \lambda W)^{-1}(\beta_{1k}I_N + \beta_{2k}W) \quad (3)$$

where  $D_k$  is the HSR connection status in spatial unit  $k$ . The total marginal effect on output that is associated with a county getting connected has two components, a direct one and an indirect one, that may both vary over space. The estimator that we use follows LeSage and Pace (2009) in reporting a single direct effect, that averages the diagonal elements of the matrix in (3) and a single indirect effect that averages the row sums of the non-diagonal elements of that matrix. Indirect effects arise not just from adjacent area units, if  $\beta_{2k} \neq 0$ , but also from (potentially) all areas through the spatial autoregressive effect if  $\lambda \neq 0$ . Thus, there can be both local and global spillovers and when these are accounted for, averages from the matrix of derivatives  $\partial O / \partial D_k$  may be quite different to the estimated direct impact effect,  $\hat{\beta}_1$ .

To further address the endogeneity of HSR route placement, we combine the IV method with equation (1). We used the straight-line strategy to construct the potential HSR connection variables, as the IV for the actual HSR connections. Compared with highways and traditional railways, early stage HSR especially aims to reduce travel time between targeted central cities. To best realize this aim, a designer might hypothetically just draw a straight line between two targeted cities. Thus, those counties and districts falling on the straight line are potential counties connected to the HSR network and this design choice can be used to construct the IV for the actual HSR connection variable. The straight-line strategy tends to produce valid IV because counties lying on the straight line of two HSR-targeted cities are more likely to have actual HSR connections, but whether or not a county is on the line is exogenous. Specifically, for each HSR line in operation, we draw a straight line between two targeted provincial capitals. Those counties on the line are potential counties for HSR connection, assigned to be in the same year the actual HSR line started to operate (see Figure 3 for an example).

Figure 3. The straight-line strategy for the instrumental variable



Note: This map illustrates the IV strategy based on potential HSR lines (shown in blue) only for 2012. The darkest areas, which take the value of 1 for the instrumental variable in the first-stage model, are the potential counties and districts connected by HSR lines based on straight lines between capital cities.

## IV. Results

### *Empirical results*

The results of estimating Eq. (1) and then imposing various restrictions on the parameters and estimating the nested models are given in Table 3 (for GDP), Table 4 (for luminosity: DMSP), and Table 5 (for luminosity: VNL). The nesting restrictions are rejected in all cases, so that the SARAR models appear to be the most data-acceptable models for both GDP and luminosity. The discussion therefore concentrates mostly on the results in column (1) for the SARAR model. Nevertheless, we also devote some discussion to the results in column (5) of each table for the standard two-way fixed effects panel data model that does not allow for any spatial lags. Those sorts of aspatial models are often used in the literature (such as some of those studies summarized in Panel A and Panel C of Table 1).

Table 3: Relationships Between HSR station opening and the Change in Economic Activity (log GDP) in China: 2012 to 2019

	(1) Spatial lag of errors, covariate and outcome	(2) Spatial lag of the covariate and outcome	(3) Spatial lag of the outcome	(4) Spatial lag of the errors	(5) Standard panel model analysis
HSR station opened	-0.005 (0.006)	-0.004 (0.006)	-0.007 (0.005)	-0.004 (0.006)	-0.025*** (0.007)
Average impacts:					
Direct	-0.004 (0.006)	-0.008 (0.006)	-0.008 (0.006)	-0.004 (0.006)	-0.025*** (0.007)
Indirect	-0.005 (0.010)	-0.052 (0.030)	-0.016 (0.012)	n.a.	n.a.
Total	-0.009 (0.011)	-0.060 (0.033)	-0.024 (0.018)	n.a.	n.a.
County fixed effects	Yes	Yes	Yes	Yes	Yes
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Spatial lag: HSRopen	Yes	Yes	No	No	No
Spatial lag: output	Yes	Yes	Yes	No	No
Spatial lag: errors	Yes	No	No	Yes	No
All covariates = 0	$\chi^2=2693^{***}$	$\chi^2=45069^{***}$	$\chi^2=45071^{***}$	$\chi^2=2780^{***}$	$\chi^2=16335^{***}$
Nesting restrictions	n.a.	$\chi^2=54935^{***}$	$\chi^2=54953^{***}$	$\chi^2=2898^{***}$	$\chi^2=70852^{***}$

*Note:* The sample period is 2012-2019, for 2342 county-level units, giving an estimation sample of n=18736. Coefficients for the fixed effects and the spatial lags are not reported. The decomposition of average impacts into direct, indirect and total components is based on LeSage and Pace (2009). The nesting restrictions are imposed on the SARAR model in column (1) to derive the models in columns (2) to (5) Standard errors are in ( ), with statistical significance at the 1%, 5% and 10% level denoted by \*\*\*, \*\*, \*.

Table 4: Relationships Between HSR station opening and the Change in Luminosity (log DMSP) in China: 2012 to 2019

	(1) Spatial lag of errors, covariate and outcome	(2) Spatial lag of the covariate and outcome	(3) Spatial lag of the outcome	(4) Spatial lag of the errors	(5) Standard panel model analysis
HSR station opened	-0.006 (0.012)	-0.003 (0.012)	-0.007 (0.011)	-0.001 (0.012)	-0.027** (0.013)
Average impacts:					
Direct	-0.001 (0.012)	-0.004 (0.012)	-0.007 (0.012)	-0.001 (0.012)	-0.027** (0.013)
Indirect	-0.035* (0.021)	-0.105** (0.044)	-0.009 (0.014)	n.a.	n.a.
Total	-0.036 (0.024)	-0.108** (0.048)	-0.016 (0.026)	n.a.	n.a.
County fixed effects	Yes	Yes	Yes	Yes	Yes
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Spatial lag: HSRopen	Yes	Yes	No	No	No
Spatial lag: output	Yes	Yes	Yes	No	No
Spatial lag: errors	Yes	No	No	Yes	No
All covariates = 0	$\chi^2=1667^{***}$	$\chi^2=21837^{***}$	$\chi^2=21829^{**}$	$\chi^2=2794^{***}$	$\chi^2=11908^{***}$
Nesting restrictions	n.a.	$\chi^2=16119^{***}$	$\chi^2=16134^{***}$	$\chi^2=1016^{***}$	$\chi^2=23655^{***}$

*Note:* The sample period is 2012-2019, for 2342 county-level units, giving an estimation sample of n=18736. Coefficients for the fixed effects and the spatial lags are not reported. The decomposition of average impacts into direct, indirect and total components is based on LeSage and Pace (2009). The nesting restrictions are imposed on the SARAR model in column (1) to derive the models in columns (2) to (5) Standard errors are in ( ), with statistical significance at the 1%, 5% and 10% level denoted by \*\*\*, \*\*, \*.

Table 5: Relationships Between HSR station opening and the Change in Luminosity (log VNL) in China: 2012 to 2019

	(1) Spatial lag of errors, covariate and outcome	(2) Spatial lag of the covariate and outcome	(3) Spatial lag of the outcome	(4) Spatial lag of the errors	(5) Standard panel model analysis
HSR station opened	-0.049*** (0.010)	-0.043*** (0.009)	-0.044*** (0.009)	-0.044*** (0.009)	-0.056*** (0.010)
Average impacts:					
Direct	-0.046*** (0.009)	-0.047*** (0.009)	-0.048*** (0.009)	-0.044*** (0.009)	-0.056*** (0.010)
Indirect	-0.019 (0.016)	-0.063* (0.034)	-0.053*** (0.011)	n.a.	n.a.
Total	-0.065*** (0.019)	-0.110** (0.037)	-0.101*** (0.020)	n.a.	n.a.
County fixed effects	Yes	Yes	Yes	Yes	Yes
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Spatial lag: HSRopen	Yes	Yes	No	No	No
Spatial lag: output	Yes	Yes	Yes	No	No
Spatial lag: errors	Yes	No	No	Yes	No
All covariates = 0	$\chi^2=1822^{***}$	$\chi^2=21950^{***}$	$\chi^2=21950^{***}$	$\chi^2=2180^{***}$	$\chi^2=12775^{***}$
Nesting restrictions	n.a.	$\chi^2=15956^{***}$	$\chi^2=15957^{***}$	$\chi^2=1090^{***}$	$\chi^2=22394^{***}$

*Note:* The sample period is 2012-2019, for 2342 county-level units, giving an estimation sample of n=18736. Coefficients for the fixed effects and the spatial lags are not reported. The decomposition of average impacts into direct, indirect and total components is based on LeSage and Pace (2009). The nesting restrictions are imposed on the SARAR model in column (1) to derive the models in columns (2) to (5) Standard errors are in ( ), with statistical significance at the 1%, 5% and 10% level denoted by \*\*\*, \*\*, \*.

The lack of data acceptability for any of the nesting restrictions indicates that for the spatial units and period we study, the interactions are occurring through the spatial lags of the outcomes, the lags of the treatment, and the lags of the errors. There are two consequences of this pattern. First, partial approaches that only allow for local spillovers, such as including the treatment status of nearby spatial units, may not reveal the full pattern of spillovers. Second, the regression coefficients by themselves do not tell the full story, and the matrices of marginal effects based on equation (3) need to be taken into consideration.

The results of the marginal effects calculations are reported in the “average impacts” rows of Tables 3, 4, and 5, using the decomposition due to LeSage and Pace (2009). These results do not indicate that HSR expansion had a positive effect on the regional economy. From column (1) of Table 3 we cannot rule out that the HSR connection has zero direct, indirect and total effect on local GDP. For the aspatial model that does not consider the spatial spillovers, column (5) of Table 3, suggests that the HSR connection is associated with post-connection local GDP being -2.5% lower compared to what it would have been based on the GDP growth of the counties without connections. The inhibiting effect on local GDP growth of a county being connected with HSR is seen very clearly with the aspatial model.

The lack of any apparent effect of HSR connections on local GDP growth should not be due to insufficient variation in the data or a research design that is somehow ‘under-powered’. The same annual GDP data, for the same spatial units, and with the same set of estimators as in Table 3, has recently been used by Zhang et al (2024) to study direct and indirect impacts on economic activity from China’s counties being administratively upgraded. In the table in that study that corresponds to Table 3 here, nine of the 11 average impact estimates were sufficiently precisely estimated to be statistically significant at  $p < 0.01$  and another was significant at the  $p < 0.05$  level. So it seems to be something more about the lack of impact of HSR connections, rather than issues with the data and estimation framework, that accounts for the imprecise estimates.

When luminosity is used as the proxy for local activity, as a sensitivity analysis in case of mistrust in China’s GDP figures, connection to the HSR network is associated with lower rates of growth in local economic activity. Using DMSP data, for the standard two-way fixed effects panel model without considering any spatial spillovers, the results in column (5) of Table 4 suggest an impact of almost minus three percent ( $p < 0.05$ ). This effect is even larger with the spatial Durbin model (in column (2)); luminosity is 11 percent lower after connection to the HSR network compared to the luminosity growth exhibited by the counties that remain

unconnected. This negative effect operates almost entirely through the indirect impacts, suggesting that there are negative spillovers from the expansion of the HSR network. The predominance of the indirect negative impacts also shows up with the SARAR results in column (1) of Table 4, but each type of impact is surrounded by wide standard errors so that indirect and total effects would only be statistically significant at  $p < 0.09$  and  $p < 0.13$ .

The clearest evidence on the negative impacts of HSR connection is with the VNL data, which are known to be far more precise and accurate than DMSP data.<sup>1</sup> The impact is almost minus six percent with the standard two-way fixed effects panel model without considering any spatial spillovers (Table 5, column (5)). In fact, across all five estimators in Table 5, the estimates of average direct impacts tightly group, varying only between -4.4% and -5.6%. Variation in estimates of total impacts comes from the indirect effects that range from a statistically insignificant -2% for the SARAR model to weakly significant -6.3% for the spatial Durbin model and a very precisely estimated -5.3% with the spatial lag model. Overall, no result in Table 5 gives support to a hypothesis of greater local economic activity after connection to the HSR network, but there is consistently strong evidence of negative effects from the expansion of HSR connections between 2012 and 2019.

### *Results from instrumental variables*

The results reported thus far are potentially biased, if the issue of endogenous placement of the HSR network matters. We therefore used an instrumental variables approach based on straight-lines which connect the major provincial cities, similar to An et al. (2022) and Liu et al. (2022). In the first stage of the two-stage least square (2SLS) results we examined the strength of this instrumental variable, with a value of the F-statistic of 1038 for omitting it from the first stage equation. Therefore, we should not have any weak instruments problem. We then repeated the models from columns (1) and (5) of Tables 3 and 5, for GDP and VNL luminosity, but using the predicted HSR connections as instruments for actual connections. The results are shown in Figure 4, with the GDP results in the top panel and the VNL results in the bottom panel.

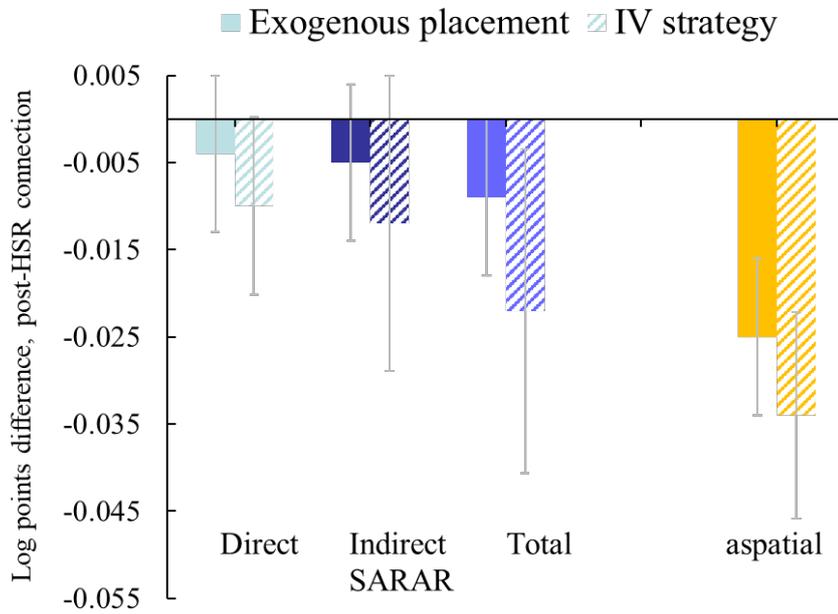
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<sup>1</sup> Prior study for China shows that county-level inequality measured with VNL data is far closer to benchmark estimates than what DMSP data show (Zhang et al, 2023) and a similar result is shown at the district-level in India (Mathen et al, 2024). In terms of evaluating treatment effects of spatially targeted interventions, Kim et al (2024) show that VNL data provide far more precise treatment effect estimates than do DMSP data.

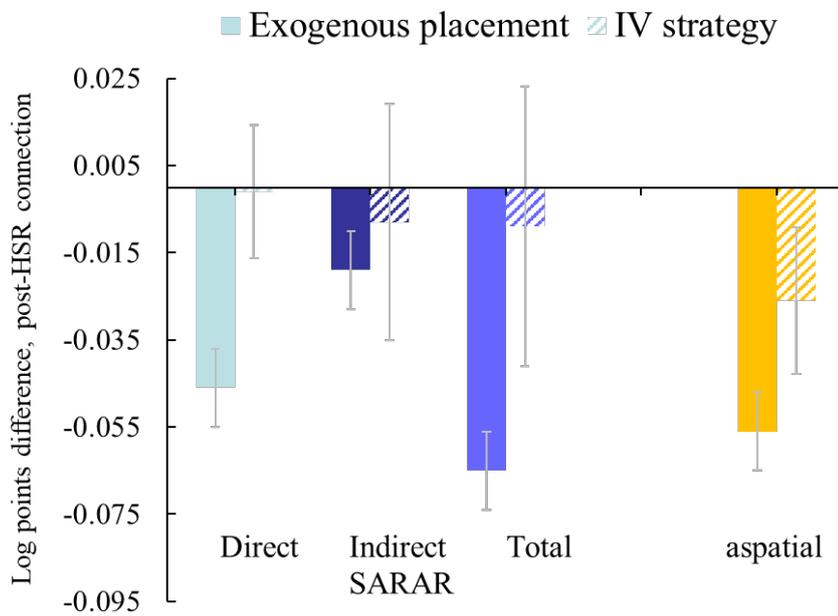
The finding of negative average total impacts on local GDP from connecting to the HSR network is amplified when the instrumental variables method is used to deal with the endogenous placement issue. The estimate of the average total impacts is more than doubled (thus, more negative) and even allowing for the larger standard errors with instrumental variables, this is a statistically significant ( $p < 0.05$ ) finding. Likewise, with aspatial two-way fixed effects, which already yielded a precisely estimated impact of -2.5% ( $p < 0.01$ ) when assuming exogenous placement, the effect when using instrumental variables becomes even more negative, at -3.4% (Fig 4a). Conversely, when luminosity ( $\log VNL$ ) is used to indicate local economic activity the use of instrumental variables causes the estimated impacts to become smaller (thus, less negative). This attenuation of the estimated negative impacts occurs with both the SARAR model and also the aspatial model. With the IV results for luminosity, one could not rule out a hypothesis of there being no effect on local economic activity from connecting to the HSR network (Fig. 4b).

**Figure 4.** Comparing HSR impacts with exogenous versus endogenous placement

**(a) log GDP as outcome variable**



**(b) log VNL luminosity as outcome variable**



Notes: Error bars show 95% CI. The IV strategy uses hypothetical straight-line connections between city pairs. Exogenous placement results are from Tables 3 and 5, column (1) and (5)

## V. Discussion and Conclusions

A dramatic expansion of high-speed rail in China presents an important empirical setting to study the impact of transport infrastructure enhancement on economic growth. Given this importance, a growing literature with dozens of studies has examined high-speed rail impacts in China. What we add to this literature is a consideration of both spatial spillovers and the endogenous placement of the high-speed rail network. Most prior studies ignore any spatial spillovers, as shown in our review of 30 papers. Indeed, we are aware of only two studies that consider both the endogeneity and the spillover issue. The other feature of our study is that we cover all parts of China, using panel data for 2342 county-level units. Many of the prior studies have focused only on a particular region, or have used more spatially aggregated data for prefectural-level cities. The growing national scope of China's high-speed rail network, and the ambition to almost double the network length over the next decade, makes a national scope for the analysis more salient.

In the period under study, the number of county-level units in China that are connected to the high-speed rail network more than tripled, with almost one-fifth of all county-level units now connected to the network. Despite this substantial increase in connectivity, our results do not show positive impacts on local economic activity. Instead, we predominantly find negative impacts, or cannot rule out nil impacts, whether allowing for spatial spillovers or not, whether allowing for endogeneity or not, and using both GDP and luminosity data as our measures of local economic activity. In this regard, our findings echo Qin (2017) and Gao et al. (2020), who note that connecting to the high-speed rail network can impede economic growth through the spatial reallocation of the economically active population from peripheral areas to more central areas, and through the spatial restructuring of industries.

Of course the priority given to investments in high-speed rail may not just be because of expected positive impacts on local GDP. The Chinese government proactively promotes indigenous technology innovation, with the high-speed rail industry a prominent example of innovation-driven development (Sun, 2015). By developing high-speed rail technology, the Chinese state successfully leapfrogged into the high-speed land transportation age, fostering innovation, and reducing dependency on foreign suppliers (Yan, 2023). These technologies serve as a fruitful starting point for China in promoting infrastructure construction in a globalized economy, such as through the Belt and Road Initiative. Moreover, even if it does

not (yet) show up in local GDP, the growth of high-speed rail may improve corporate financing efficiency and optimise the efficiency of resource allocation in capital markets (Jin, Zhang, and Xin, 2020). For example, being connected to the network significantly increases a city's capital mobility (Duan et al. 2021), and the resource redistribution between regions (Meng, Lin, and Zhu 2018).

Notwithstanding these considerations, the apparent negative impacts on county-level economic growth, and potentially negative spillover effects on neighbouring spatial units does pose a problem as China adapts to lower expectations of future economic growth. Before planning and constructing future additions to the high-speed rail network, and future stations on the network, policy makers may need to do more thorough evaluation of economic impacts on the targeted county-level units and also on their neighbours.

Future research could delve into the heterogeneous effects of high-speed rail on economic growth and explore broader impacts on the environment and on society. For example, some studies have proposed that impacts of high-speed rail spatially vary along the route, tending to benefit developed cities with several rail stations and higher accessibility rather than least-developed and developing cities with remote locations and backward economies (Liang et al., 2020; Zou and Chen, 2024). Furthermore, any spillover effects may be examined using a broader set of weighting matrices to what we have used, such as allowing for asymmetric effects that reflect greater agglomeration forces in some directions rather than others, or in some regions rather than others (as discussed, for example, in Zhang et al, 2024). These potential refinements provide a direction and focus for our future research.

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